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On the reliability of the Elements of Metacommunity Structure framework for separating idealized metacommunity patterns

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Abstract

The Elements of Metacommunity Structure (EMS) framework originally suggested by Leibold and Mikkelsen (2002) in *Oikos* is a popular approach to identify idealized metacommunity patterns (i.e. checkerboard, nested, evenly spaced, Clementsian, Gleasonian), and hereby to infer the existence of structuring processes in metacommunities. Essentially, the EMS framework consists of the rearrangement of the sites-by-species incidence matrix followed by a series of tests for coherence, turnover and boundary clumping in species distributions. Here, we give a critical evaluation of the EMS framework based on theoretical considerations and simulations. We found that user defined site ordering may influence the coherence test (number of embedded absences) depending also on the ordering of species, and therefore we argue that the application of user-defined matrix rearrangement has strong limitations. The recommended ordering by correspondence analysis is sensitive to matrix structure and may even include arbitrary decisions in special cases. Further, we revealed different meanings of the checkerboard pattern and showed that negative coherence is not necessarily associated with this as assumed in the EMS framework. Also, the turnover test cannot always detect nested pattern, because turnover and nestedness are not necessarily the opposite endpoints of a continuum. We argue that the boundary clumping test can only be used for separating Clementsian, Gleasonian and evenly spaced patterns if sites are ordered along a real environmental gradient rather than a latent one identified by correspondence analysis. We found that the series of tests in the EMS framework are burdened by anomalies and that the detection of some metacommunity patterns is sensitive to type II error. In sum, our findings suggest that the analytical methodology of the EMS framework, as well as the conclusions drawn from its application to metacommunity studies require careful reconsideration.

44

Keywords

biodiversity; community pattern; pattern analysis; idealized metacommunity patterns

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49 **1. Introduction**

50 Detecting and understanding drivers of metacommunity structure are key issues in
51 community ecology with significant legacy (Mittelbach 2012). Early ecologists have already
52 inferred the existence of structuring forces from the community patterns observed. For
53 instance, Clements (1916), the pioneer of North American plant ecology, viewed plant
54 communities as coherent units with discrete boundaries formed in response to
55 environmental factors (*Clementsian pattern*). In contrast, Gleason (1926) argued that species
56 have distinct ecological characteristics and therefore individualistic responses to underlying
57 environmental gradients (*Gleasonian pattern*). *Evenly spaced pattern* occurs in systems with
58 trade-offs in fitness in different environments, resulting in a spatial distribution with evenly
59 dispersed populations (Tilman 1982). Intense interspecific competition may generate
60 *checkerboard pattern* where pairs of species are mutually exclusive (Diamond 1975). Finally,
61 *nested pattern* occurs when species poor communities consist of subsets of species
62 occurring in richer communities (Patterson and Atmar 1989). These cases have been
63 regarded as idealized types of metacommunity pattern (Ulrich and Gotelli 2013, Heino et al.
64 2015) and have received increasing attention due to their theoretical interpretation
65 (Carvalho et al. 2013, Ulrich and Gotelli 2013).

66 The development of metacommunity theory provided a conceptual framework for ecologists
67 to disentangle underlying drivers (niche based species sorting, dispersal, drift, see Vellend
68 2010, Shipley et al. 2012) of multisite communities (Leibold et al. 2004). Some of the
69 approaches use multispecies distribution patterns for inferring the existence of structuring
70 ecological forces. No doubt that the “elements of metacommunity structure” approach
71 suggested by Leibold and Mikkelsen (2002) and its upgrade (Presley et al 2010, hereafter
72 referred to as EMS framework) provide a very popular methodology developed for this
73 purpose.

74 The EMS framework includes the rearrangement of the sites-by-species incidence matrix
75 followed by three tests (Fig. 1), each related to a given element of metacommunity
76 structure. First, the rows and the columns of the matrix are ordered along the first axis of
77 correspondence analysis (CoA) to discern variation in response to a latent environmental
78 gradient. According to Leibold and Mikkelsen (2002, p. 241), the simultaneous ordering of
79 sites and species has three purposes: (1) it often minimizes the number of interruptions in
80 species' ranges (number of embedded absences), (2) it provides a basis for judging whether
81 a given metacommunity is nested, or dominated by turnover (high number of species
82 replacements), and (3) it defines the boundaries of species' ranges (boundary clumping).
83 Consequently, matrix rearrangement via CoA has strong impact on the assessment of each
84 element of metacommunity structure. Note that although this procedure is recommended
85 for general use, the EMS framework also allows user-defined matrix ordering. Secondly
86 coherence, the first element of metacommunity structure is defined as the number of
87 embedded absences in the matrix and its significance is examined using null model tests.
88 Following the study of Gotelli (2000), species richness of sites is kept constant in the
89 recommended null model (Presley et al. 2010). If coherence is negative (the number of
90 embedded absences is significantly higher than expected by chance) then the EMS

91 framework detects checkerboard pattern. If the number of embedded absences does not
92 differ significantly from a randomly generated value (coherence is random) then the EMS
93 framework indicates a random pattern. If coherence is positive (the number of embedded
94 absences is lower than expected by chance) then the matrix should be examined for
95 turnover. Turnover, the second element of metacommunity structure, is measured as the
96 number of times one species replaces another between two sites (i.e. number of
97 replacements) for each possible pair of species and for each possible pair of sites. If turnover
98 is negative (the number of replacements is lower than expected by chance) then the EMS
99 framework reveals a nested pattern, if turnover is random the EMS detects quasi pattern
100 (see Presley et al. 2010), and if turnover is positive (the number of replacements is higher
101 than expected by chance) then the EMS framework suggests the existence of Clementsian,
102 Gleasonian or evenly spaced patterns. These latter three are separated from each other by
103 examining the boundary clumping of species ranges, the third element of metacommunity
104 structure, using the Morisita test. If clumping is positive (Morisita I is significantly larger than
105 1.0) then the EMS framework detects Clementsian pattern; if clumping is negative (Morisita I
106 is significantly lower than 1.0) evenly spaced pattern is indicated, and if clumping is random
107 (Morisita I does not significantly differ from 1.0) then the pattern is thought to be
108 Gleasonian.

109 There is, however, much controversy about the relative merits of the EMS framework.
110 Gotelli and Ulrich (2012, p. 178), for instance, noted that species segregation and
111 aggregation examined in the coherence test "might be the different sides of the same coin"
112 and that rearranging the matrix (i.e. the reordering of sites by correspondence analysis)
113 "does not alter any of the underlying information on species occurrences in the matrix". By
114 examining the power of different null model algorithms, Gotelli and Ulrich (2012) found that
115 a segregation measure was not exactly opposite in its behavior to a nestedness measure,
116 suggesting that nested and segregated patterns (i.e. evenly spaced, Gleasonian and
117 Clementsian) are not necessarily mutually exclusive as implied by the turnover test in the
118 EMS framework. The same authors repeated this comment later and also argued that "The
119 frameworks proposed by Leibold and Mikkelsen (2002), and Presley et al. (2010) implicitly
120 assume that measures of coherence, turnover, and boundary clumping describe orthogonal,
121 independent properties of matrices. But if the measures are strongly correlated, some of the
122 proposed cells in their classification frameworks may be redundant or not achievable.
123 Leibold and Mikkelsen (2002) recognized this problem and noted that they were able to
124 identify empirical matrices that fit each of the five different scenarios they described" (Ulrich
125 and Gotelli 2013, p. 3). A more recent paper stated that the efficiency of the EMS framework
126 is heavily dependent on data quality (Mihaljevic et al. 2015, see also Gotelli and Graves
127 1996, Ulrich and Gotelli 2013) and suggested the use of occupancy models to at least partly
128 overcome this problem. These models allow an estimation of predicted occupancy at each
129 sample site and thus make it possible to distinguish between the probability of a species
130 occurring at a site and the probability of a species being detected at a site in which it does
131 occur (Mihaljevic et al. 2015). These critical comments, however, did not prevent community
132 ecologists from using the methodology even further. The EMS framework has still been used
133 increasingly both in terrestrial and aquatic realms for finding the best fit to idealized

134 metacommunity patterns (Dallas and Presley 2014, de la Sancha 2014, Heino et al. 2015).
135 However, the reliability of the method in discerning idealized (meta)community patterns has
136 not been tested as yet.

137 To fill this methodological gap, this paper examines the performance of the EMS framework.
138 Combining theoretical aspects with simulation approaches we go through this approach step
139 by step and inspect how the rearrangement of the matrix, the output of individual tests as
140 well as their series influence the success of analysis. We examined also the robustness of the
141 methodology to increasing noise in the data, as well as the practice of researchers in
142 revealing the importance of environmental factors structuring metacommunity patterns.

143

144 **2. Methods**

145 To guarantee unambiguous answers, we first carefully review terms and procedures related
146 to the EMS framework. We discuss possible interpretations of terms and evaluate the
147 performance of different procedures. In case of equivocal use of any term or procedure, we
148 attempt to clarify the situation by suggesting a solution.

149 We calculated the following indices: the number of embedded absences (the index of
150 coherence test, Leibold and Mikkelsen 2002, Presley et al. 2010), the number of mutually
151 exclusive species pairs (Diamond 1975), turnover (the index of turnover test, Leibold and
152 Mikkelsen 2002, Presley et al. 2010). As nestedness is not defined in the EMS framework, we
153 used two nestedness measures, the relativized nestedness (N_{rel} , Podani and Schmera 2011)
154 and the site-order independent version of NODF (Almeida-Neto et al. 2008) called as
155 $NODF_{max}$ (Podani and Schmera 2012, Ulrich and Almeida-Neto 2012).

156 We examined the behavior of indices themselves as well as the behavior of the indices in
157 null model tests. Indices were examined using toy data sets in series of site-by-species
158 incidence matrices. We examined the relationship between indices in two-site situations
159 using the random parameter approach (Chao et al. 2012, see also Baselga and Leprieur
160 2015). In the first (Random parameter approach 1), we assumed that the numbers of species
161 present in both sites (a), present only in the first site (b), and present only in the second site
162 (c) are derived from a uniform distribution ranging from 0 to 100. We generated 50,000
163 triplets of random a , b and c values, and removed data records with empty sites. In the
164 second case (Random parameter approach 2), we assumed that 200 species are distributed
165 among the three sets (a , b and c). We produced all possible combinations and removed data
166 records with empty sites. Furthermore, we simulated all the possible sites-by-species
167 matrices containing 4 sites and 4 species (degenerate matrices were omitted). This
168 procedure resulted in 41,503 binary matrices, called hereafter as 4-by-4 binary matrices.

169 Although the 4-by-4 binary matrices allow examining the response of indices to all possible
170 situations in the matrix, the null model test of the matrix might be problematic due to the
171 small number of sites and species. We therefore produced 10,000 random matrices with 10
172 sites and 10 species (degenerate matrices, i.e. those containing empty rows or columns,
173 were omitted). These are referred to as 10-by-10 matrices. We used them in null model tests

174 (Gotelli and Graves 1996). For each random matrix, we generated 1000 null matrices.
175 Although there are many algorithms to produce 'random' or 'null' matrices and these
176 algorithms have different statistical properties and ecological meanings (Gotelli and Ulrich
177 2012, Ulrich and Gotelli 2013, Strona et al. 2017), we selected the null model method that
178 maintained the species richness of every site and filled species ranges based on their
179 marginal probabilities ("r1" method in *metacom* package, Dallas 2014). The P value
180 (estimated probability of type I error) was calculated as the number of null matrices whose
181 index value was more extreme than or equal to the observed index. We applied a two-tailed
182 test at $\alpha = 0.05$. The Jaccard index (Jaccard 1912) was used to measure the similarity of
183 different null model tests: the number of matrices proved to be significantly positive (or
184 negative) in both tests was divided by the number of such matrices plus those that were
185 found significant only in either of the two tests. Positive and negative results in the two tests
186 were not distinguished.

187 We used a noise test (Gotelli 2000, Podani and Schmera 2012) to examine the sensitivity of
188 the EMS framework to increasing randomness in community data. We started with 20-by-20
189 perfectly structured nested, Gleasonian, evenly spaced and Clementsian patterns (Electronic
190 Appendix 1). These patterns were regarded as initial patterns (step 0, 0% noise). We then
191 gradually added noise (randomness) to the matrix in the following way: In the first step (5%
192 noise), 20 pairs of randomly chosen values in the matrix were interchanged (referred to the
193 full randomization model in Podani and Schmera 2012). In the second step (10% noise), 40
194 pairs of randomly chosen values were interchanged. Complete randomness (100% noise) is
195 achieved after 20 steps, with a total of 400 interchanges. Degenerate matrices were
196 omitted. This procedure was repeated 100 times for every step. EMS analysis was performed
197 for each step (21 steps) 100 times. The output of the noise test shows the relative frequency
198 of detected metacommunity patterns in response to increasing noise level (from 0% to
199 100%). The ideal - and expected - situation is that at low noise level the methodology detects
200 mostly the initial pattern. At intermediate noise level, the initial pattern is detected in a
201 decreasing number of times, while the frequency of random pattern is increasing. At high
202 noise level, the frequency of random pattern should be the largest. If the initial pattern is
203 not detected many times even at low noise level, then the EMS framework is sensitive to
204 type II error. In contrast, if the initial pattern is detected with high frequency even at high
205 noise level, the EMS framework is sensitive to type I error.

206 Finally, we examined how researchers use the EMS framework and handle the importance of
207 environmental factors in shaping metacommunity patterns. To reveal this, first we made a
208 search using ISI Web of Science (access date: 28 July 2015) on the number of papers citing
209 Presley et al. (2010). In the second step, we searched for papers applying the EMS
210 framework. We divided these papers into two groups: those applying user defined matrix
211 ordering and articles using CoA for site and species ordering. Then, we searched for papers
212 that reported the variance explained by CoA axes. In our view, this information is essential,
213 and should be obligatorily added to EMS analysis as an expression of the reliability of the
214 method. No doubt that the amount of community variation explained must be used for

215 assigning the studied metacommunity to an idealized pattern. Finally, we examined whether
216 the axes of CoA (EMS framework) were related to any environmental variables.

217 All calculations were performed in R (R Core Team 2016). All possible matrices containing 4
218 sites and 4 species were produced by the *gtools* package (Warnes et al. 2015). Null matrices
219 were produced by the *metacom* package (Dallas 2014). Correspondence analysis (CoA) was
220 performed by the *ca* package (Nenadic and Greenacre 2007), the number of mutually
221 exclusive species pairs, number of embedded absences, turnover, relativized nestedness and
222 $NODF_{\max}$ were calculated by R-scripts developed by the authors (Electronic Appendix 2).

223

224 **3. Site and species ordering**

225 By definition, site and species orderings influence the number of embedded absences
226 (order-dependent measure) in the data matrix, but they have no impact on the number of
227 replacements (order-independent measure). In addition, site ordering also affects patterns
228 in boundary clumping (order-dependent measure). That site and species ordering both
229 influence coherence can be explained by the definition of embedded absence: "an
230 interruption in a range or community" (p. 242 in Leibold and Mikkelsen 2002).

231 Studying communities along an environmental gradient is a typical situation for user-defined
232 site-ordering. The EMS framework allows user-defined matrix ordering without emphasizing
233 the importance of species ordering. Since coherence is influenced not only by the order of
234 sites but also by species ordering, as said, user-defined matrix ordering has strong
235 limitations. Therefore, if the data matrix is ordered by the user, we recommend a clear
236 definition of species ordering, if it is possible at all.

237 Alternatively, the recommended matrix-ordering uses the first axis of CoA to define the
238 order of sites (and species) for the coherence test. In this case, we disclaim real
239 environmental gradients and focus on the "within-matrix data structure". In complex data
240 structures, however, the first axis of CoA does not necessarily explain considerably more
241 variation than the subsequent axes. In other words, the first axis of CoA might identify *one*
242 *dominant* but not necessarily *the only dominant* axis of community variation. This means
243 that analyses of the same data matrix reordered along different axes might reveal
244 contrasting aspects of data structure. We by no means state that the use of the first axis of
245 CoA is a bad decision but emphasize that further studies are needed to reveal the effect of
246 choosing among similarly important axes, and to merge alternative results into a consensus.

247 Finally, although CoA has been one of the most popular ordination methods in numerical
248 ecology, it has some limitations. One of these is that sites with single and unique species
249 cannot be ordered due to the lack of overlap with other sites and species. If software
250 packages do order such matrices "in a way" then the result is based on an *arbitrary decision*
251 (Electronic Appendix 3). Accordingly, CoA performed by different software packages may
252 provide differently ordered matrices and thus it cannot be regarded as "a standardized
253 approach to order sites and species" as stated by Presley et al. (2010, p. 910). Although field
254 ecologists might argue that actual data sets rarely contain sites with unique species, we have

255 three arguments in favor of discussing this situation. First, a methodology should work under
 256 all circumstances, or at least its users should be aware of any limitations. Second, the chance
 257 of observing sites with unique species cannot be excluded completely in actual data sets.
 258 Finally, checkerboard pattern, a key term of the EMS framework, has a strong theoretical
 259 connection to sites with unique species (see next paragraph).

260

261 **4. The multiple meaning of checkerboard pattern**

262 Since the coherence test of the EMS framework is supposed to separate checkerboard,
 263 random and other data structures from each other, first we review the meaning of
 264 checkerboard pattern and then identify its most conspicuous realization. In examining the
 265 co-occurrence of bird species on islands, Diamond (1975) proposed the term "checkerboard
 266 distribution" for competing pairs of species with mutually exclusive island-by-island
 267 distributions. In this metaphor, the distribution of the competing pair of species reflects the
 268 alternating squares of dark and light colors in a checkerboard pattern (of which the
 269 chessboard is a special case). Accordingly, we can visualize this checkerboard distribution for
 270 a given pair of species by the following matrix (\mathbf{M}_1), where sites are rows and species are
 271 columns:

272

$$273 \quad \mathbf{M}_1 = \begin{bmatrix} 1 & 0 \\ 0 & 1 \\ 1 & 0 \\ 0 & 1 \end{bmatrix}$$

274

275 It is very important to emphasize that Diamond's original checkerboard distribution reflects
 276 the distribution of pairs of species, and the "checkerboard character" of a community has
 277 been mostly (but not always, see Stone and Roberts 1990, 1992; Gotelli 2000) characterized
 278 by the *number of checkerboard species pairs* (i.e. the number of species pairs with non-
 279 overlapping occurrence). In agreement with this, and for compatibility with the EMS
 280 framework (Presley et al. 2010), we will use the term *checkerboard pattern* for binary
 281 matrices in which the number of checkerboard species pairs (Stone and Roberts 1990) is
 282 high. This definition means that the checkerboard pattern may be identified in binary
 283 (presence-absence) matrices without any restriction as to the order of sites and species.

284 Almeida-Neto et al. (2008) used the term "checkerboard" for any binary matrix in which
 285 every cell containing the value of 1 has the same value in all diagonally neighboring cells and
 286 0 in the remaining neighboring cells, for example:

287

288 $\mathbf{M}_2 = \begin{bmatrix} 1 & 0 & 1 & 0 \\ 0 & 1 & 0 & 1 \\ 1 & 0 & 1 & 0 \\ 0 & 1 & 0 & 1 \end{bmatrix}$

289

290 We refer to this as *visual checkerboard pattern*, which has been used mostly for visualization
 291 purposes. Since its first appearance (Almeida-Neto et al. 2008), the visual checkerboard
 292 pattern has been used frequently as an example matrix for developing pattern analysis
 293 procedures (Podani and Schmera 2011, Ulrich and Gotelli 2013). However, many analytical
 294 tools in metacommunity ecology are insensitive to the order of sites and species in the
 295 matrix and therefore are unable to detect and test this unique property. A noted example is
 296 the compartment pattern, which differs from the visual checkerboard pattern only in the
 297 ordering of sites and species (Podani and Schmera 2011, Ulrich and Gotelli 2013).

298 Recently, Connor et al. (2013) have contributed by two very important points to the proper
 299 interpretation of the checkerboard pattern in situations where ordering is fixed by
 300 geographical constraints. First, they argued that the checkerboard metaphor reflects the
 301 mutually exclusive distribution of two species on a set of islands, where only one of the two
 302 species (denoted by the letters A or B) is present on a single island, while the position of the
 303 letters corresponds to the explicit geographic position of the islands. Consequently, a set of
 304 16 islands arranged spatially in a regular 4-by-4 grid and occupied by two species (A and B)
 305 shown below indicates a checkerboard pattern:

306

307 $\begin{bmatrix} A & B & A & B \\ B & A & B & A \\ A & B & A & B \\ B & A & B & A \end{bmatrix}$

308

309 Connor et al. (2013) argued also that the metaphor does not intend to visualize the
 310 presence-absence of the species in a sites-by-species incidence matrix (for example, \mathbf{M}_2 as
 311 given above) and thus this incidence matrix, in disagreement with other studies (Almeida-
 312 Neto et al. 2008, Podani and Schmera 2011, Ulrich and Gotelli 2013), should not be regarded
 313 as a "real" spatial checkerboard pattern.

314 As a second contribution, Connor et al. (2013) pointed out that the idea of Diamond's
 315 checkerboard distribution includes not only mutually exclusive island-by-island distribution
 316 of species pairs, but also the overlapping geographic ranges of the species. According to this
 317 argument and without any information on the distribution ranges of species, the lack of co-
 318 occurrence due to competition cannot be separated from lack of co-occurrence owing to
 319 non-overlapping ranges (spatial turnover). All of these suggest that presence-absence

320 matrices fail to address Diamond's (1975) original idea because a binary data matrix by itself
321 has "no explicit geography" and provides no information on the distribution ranges of
322 species. To clarify the situation, Connor et al. (2013) suggested the term *true checkerboard*
323 *pattern* for a pair of species which never co-occur on the same island and the islands
324 occupied by these two species are geographically alternating.

325 In sum, we will use the term *checkerboard pattern* when the binary matrix contains a large
326 number of checkerboard species pairs. *Visual checkerboard pattern* refers to a binary matrix,
327 in which zeros regularly alternate with 1-s over columns and rows. Finally, *true checkerboard*
328 *pattern* refers to a pair of species which never co-occur on the same island and the islands
329 occupied by these two species geographically alternate.

330 The EMS framework (Presley et al. 2010) refers to Diamond's definition (Diamond 1975), and
331 thus to *checkerboard pattern*. Some studies using the EMS framework, however, apparently
332 have to do with the *visual checkerboard pattern* (see Fig. 2 in Tonkin et al. 2017) to which
333 this methodology does not apply.

334 Of the 41,503 4-by-4 binary matrices, the largest number of mutually exclusive species pairs
335 (6) was observed when all sites contained only a unique species (Electronic Appendix 4). This
336 suggests that if the checkerboard pattern is quantified by the number of checkerboard
337 species pairs (as in the EMS framework), then sites with single and unique species will
338 contribute the most to the checkerboard character. Unfortunately, ordering of such matrices
339 is arbitrary in correspondence analysis (Electronic Appendix 3), and thus the application of
340 CoA and exclusive species pairs within the same approach may not be optimal.

341

342 **5. Coherence test**

343 The first promise of the coherence test is that a high number of embedded absences
344 (negative coherence) indicates checkerboard pattern (high number of checkerboard species
345 pairs). The examination of 4-by-4 binary matrices shows that sites with single and unique
346 species (matrices with the highest number of checkerboard species pairs) have no
347 embedded absences. To get a deeper insight into this relationship, we plotted the number of
348 embedded absences (quantifying negative coherence) against the number of checkerboard
349 species pairs (quantifying checkerboard pattern) for the 41,503 4-by-4 incidence matrices
350 (Fig. 2). As seen, a high number of embedded absences is not necessarily associated with low
351 number of checkerboard species pairs and vice versa. Moreover, 4-by-4 matrices with the
352 highest number of checkerboard species pairs (sites with single and unique species) contain
353 no embedded absences, while matrices with the largest number of embedded absences (not
354 shown) do not contain checkerboard species pairs.

355 The analyses of 10-by-10 matrices revealed that 303 matrices showed a significantly higher
356 number of embedded absences (negative coherence) than expected and thus exhibited
357 checkerboard pattern. The null model test detected 66 matrices with significantly large
358 number of checkerboard species pairs, of which 15 matrices were selected also by the
359 coherence test. This suggests that 4.24% (Jaccard similarity = 15/354) is the agreement

360 between null model tests using the number of embedded absences (coherence test of the
 361 EMS framework) and the number of checkerboard species pairs. Consequently, the number
 362 of embedded absences does not necessarily indicate checkerboard pattern and thus cannot
 363 be used alone as its indicator.

364

365 **6. Definition of turnover and its test**

366 The EMS framework assumes that turnover and nestedness are opposing patterns (Leibold
 367 and Mikkelsen 2002, p. 239). This means that if we observe low nestedness then turnover
 368 should be high and vice versa. In an extreme situation, zero nestedness should yield
 369 maximum turnover. To illuminate the relationship between turnover and nestedness, let us
 370 examine the following example (rows are sites and species are columns):

371

372 1000000000

373 0111111111

374

375 Here turnover (number of times one species replaces another between two sites: in a two-
 376 site situation it means $b*c$, where b is the number of species present only in the first, while c
 377 is the number of species present only in the second site, Presley et al. 2010) equals to 9.
 378 Note that in a 2-by-10 incidence matrix the maximum possible value of turnover is 25
 379 (obtained when $b = 5$ and $c = 5$). Consequently, this turnover value is 64% lower than the
 380 theoretical maximum. Although several nestedness indices do exist (the EMS framework
 381 does not define any measure), all of them agree that if two sites do not share any species
 382 then nestedness should be zero (Ulrich et al. 2009, Podani and Schmera 2012). Thus, this
 383 example demonstrates a situation with relatively low turnover and zero nestedness.

384 Before discussing the relationship between turnover and nestedness, we should note that
 385 the turnover definition applied by Presley et al. (2010) and used here differs from many
 386 existing definitions of turnover (see Tuomisto 2010, Anderson et al. 2011, Gotelli and Ulrich
 387 2012). We by no means state that this measure does not quantify the concept of turnover,
 388 but emphasize its uniqueness in community ecology and therefore further studies are
 389 needed to clarify its performance.

390 We examined the relationship between turnover and nestedness in two-site situations using
 391 both random parameter approaches (Fig. 3). All combinations of nestedness measures and
 392 random parameter approaches showed that high turnover associates mostly with low
 393 nestedness. However, low turnover values can be associated with a wide range of
 394 nestedness values, suggesting that turnover and nestedness are not necessarily opposing
 395 patterns. Although under specific conditions we can assume that high turnover predicts low
 396 nestedness, this is not always the case (see Random parameter approach 2). On the other
 397 hand, low turnover does not necessarily predict high nestedness.

398 We studied the relationship between turnover and nestedness using all possible 4-by-4
 399 matrices. When nestedness was quantified by the relativized nestedness measure, we found
 400 a relatively strong negative relationship between the two variables ($r = -0.860$, Fig. 4).
 401 Although low turnover values indicate high relative nestedness, high turnover does not
 402 necessarily indicate low relativized nestedness. When nestedness was quantified by
 403 $NODF_{max}$, the negative relationship with turnover was lower than with relativized nestedness
 404 ($r = -0.641$, Fig 5), and a low turnover value may be indicative of low nestedness.

405 We used null model tests on 10-by-10 matrices to examine whether significantly high
 406 turnover is associated with significantly low nestedness, and whether significantly low
 407 turnover with high nestedness. Null model tests indicated 421 matrices with high turnover
 408 and 433 matrices with low nestedness when the latter is measured by the relativized
 409 measure. The agreement between the two assessments was 23.59% (i.e. Jaccard similarity =
 410 $163/691$). When nestedness was quantified by $NODF_{max}$, 296 matrices showed low
 411 nestedness. The agreement between high turnover and nestedness ($NODF_{max}$) was only
 412 5.60% (Jaccard similarity = $38/679$). None of our null model tests indicated significantly low
 413 turnover, high relativized nestedness and high $NODF_{max}$. These results suggest that high
 414 turnover is not necessarily associated with low nestedness in the null model tests. In
 415 agreement with these findings, Ulrich and Gotelli (2013) and Ulrich et al. (2017) have already
 416 published similar results.

417

418 **7. Boundary clumping test**

419 Our starting point is that Clementsian, Gleasonian and evenly spaced patterns can only be
 420 interpreted along an actual (real) environmental gradient (Clements, 1916, Gleason 1926,
 421 Tilman 1982, Shipley and Keddy 1987). We argue that "within matrix data structure"
 422 revealed by CoA is inadequate for this purpose. Although user defined site ordering might
 423 allow testing real environmental gradients, difficulties associated with the coherence test
 424 (number of embedded absences is influenced by species ordering, unclear interpretation of
 425 coherence) strongly limit this possibility. All of these suggest that no boundary clumping test
 426 can be performed within the context of the EMS framework.

427

428 **8. Series of tests**

429 The EMS framework includes a well-defined sequence of three tests (coherence, turnover
 430 and boundary clumping). If we assume that these tests indicate orthogonal and independent
 431 properties of matrices then all these tests could be performed independently from the
 432 results of tests made earlier in the series. Although Leibold and Mikkelsen (2002, p. 239)
 433 argue that "turnover and clumping are most meaningful in the context of reasonably
 434 coherent ranges", the application of a series of tests has strong consequences. First, a test
 435 performed in a series or alone has different statistical and ecological meanings. For instance,
 436 the turnover test alone indicates the concept where species are replaced by one another,
 437 while within the EMS framework it indicates the existence of replacement in positively

438 coherent metacommunity patterns. Second, some patterns should be more frequently
 439 detected than others because earlier tests restrict the possible outputs (i.e. turnover test
 440 can only be performed if coherence is high and cannot be performed when coherence is
 441 random or negative). Consequently, if we perform a two-tailed statistical test with 5%
 442 significance level, then about 2.5% of the examined random matrices should show
 443 checkerboard pattern, 95% random pattern, 0.0625% ($2.5\% \times 2.5\%$) nested pattern, 2.375%
 444 ($2.5\% \times 95\%$) quasi pattern, 0.0015625% ($2.5\% \times 2.5\% \times 2.5\%$) evenly spaced and
 445 Clementsian pattern, and 0.059375% ($2.5\% \times 2.5\% \times 95\%$) Gleasonian pattern in a series of
 446 tests suggested by the EMS framework.

447 If we assume that the tests are not orthogonal and not independent then a series of tests
 448 may have a clear ecological meaning. In this case, however, the output of an earlier test
 449 should predict the output of a later test, or the ecological meaning of the output of an
 450 earlier test suggests that there is no need for further ecological information. The argument
 451 of Leibold and Mikkelsen (2002, p. 239) that "turnover and clumping are most meaningful
 452 in the context of reasonably coherent ranges" suggests that the EMS framework considers
 453 coherence as primary feature of metacommunity organization. However, we see no strong
 454 theoretical support for the priority of coherence in metacommunity structuring.

455

456 **9. Noise test**

457 The noise test showed that the reliability of the method to identify idealized structures is
 458 different at the same level of noise (Fig. 6, Electronic Appendix 5). For example, the noise
 459 level at which the idealized structure was detected at least with 50% reliability was below
 460 only 5%, 5%, 10% and 20% for the evenly spaced, Gleasonian, nested and Clementsian
 461 gradients, respectively. These results may explain why Clementsian (or quasi-Clementsian)
 462 and nested patterns are identified most frequently in actual data sets and why Gleasonian
 463 pattern is identified relatively infrequently. Further, identification of the evenly spaced
 464 gradient was not possible in all cases even at zero noise. As low as 5% noise in the data
 465 already yielded that the EMS method identified either Clementsian or Gleasonian structure.
 466 Above 15% (nested), 20% (evenly spaced), 25% (Gleasonian) or 30% (Clementsian) noise
 467 levels, the EMS method identified random metacommunity structure in more than 50% of
 468 cases, which further proves the sensitivity of the method to the characterization of idealized
 469 structures at different noise levels.

470

471 **10. The importance of environmental factors in shaping metacommunity patterns: current** 472 **practice**

473 We found fifty papers citing Presley et al. (2010). Twenty-six papers, each of them ordering
 474 sites and species by correspondence analysis, applied the EMS framework. We found that
 475 only four papers out of these 26 attempted to provide information on the variance explained
 476 by CoA in some way. Three of these 4 papers provided eigenvalues of the first two axes of
 477 CoA. However, these two eigenvalues by themselves do not quantify the percentage of

478 variance they explained. There was a single paper of the 26 (3.8%) that provided information
479 on the variance explained by CoA. This paper showed also that the first axis of CoA
480 accounted for 17.7% to 24.0% community variation depending on the metacommunity
481 studied and that environmental variables explained 47.9% to 77.4% variance of the first axis
482 of CoA (Erős et al. 2014). Although this single study does not allow general conclusions to be
483 drawn, it implies that metacommunity patterns are detected based on a limited amount of
484 community variation and that this limited community variation is correlated only at an
485 intermediate-level with multiple environmental variables. In most studies, the variance
486 explained by CoA is not given at all, only the relationship between the site position in
487 ordination axis and environmental variables (80.8%). In some cases, CoA is used in the EMS
488 framework, but environmental variables are related to canonical correspondence analysis by
489 the reasoning that canonical correspondence analysis is related to CoA (de la Sancha et al.
490 2014), or can be regarded as a constrained extension of CoA (Heino et al. 2015). These
491 studies ignore the fact that CoA and canonical correspondence analysis need not result in
492 the same ordering of sites along any ordination axis. Overall, our literature survey shows
493 that essential information, including at least some hints on the reliability of the identification
494 of idealized metacommunity structures remains completely hidden in almost all studies
495 which used the EMS framework.

496

497 **11. Conclusions**

498 Our theoretical and statistical considerations show that the EMS framework has to be used
499 with caution for the identification of idealized metacommunity patterns. While it is
500 appealing to identify the best-fit metacommunity structure under a single analytical
501 framework, the reliability of the test to distinguish among the idealized structures is strongly
502 case dependent.

503 We showed that although user-defined site-ordering allows testing the response of
504 community to an actual environmental gradient, its application is problematic due to the
505 dependence of the coherence test upon the order of species. Unfortunately, this
506 dependence strongly limits the performance of the EMS framework in testing the response
507 of communities to real environmental gradients. Even if CoA is used for the ordering of sites,
508 the EMS framework is relatively unreliable for separating evenly spaced, Gleasonian and
509 Clementsian patterns. Our results demonstrate that the coherence test is the most critical
510 step of the EMS framework. We found that it is not necessarily adequate for separating
511 checkerboard pattern and its ecological meaning is not clearly defined. Our observations are
512 in strong agreement with the findings of Gotelli and Ulrich (2012) in that the turnover test is
513 not necessarily adequate for detecting a nested pattern.

514 We concluded that the application of a series of tests requires further considerations and
515 that the detection of some idealized patterns is prone to type II error. Our literature survey
516 clearly indicated that the documentation of the results of the EMS framework analysis is
517 insufficient and thus information is extremely limited on the amount of community variation
518 used for detecting idealized metacommunity patterns and also on the relationship between

519 this variation and environmental drivers. These findings call for reconsidering the analytical
520 steps of the EMS framework, and for careful interpretation of its results.

521

522 **Acknowledgements**

523

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525

526

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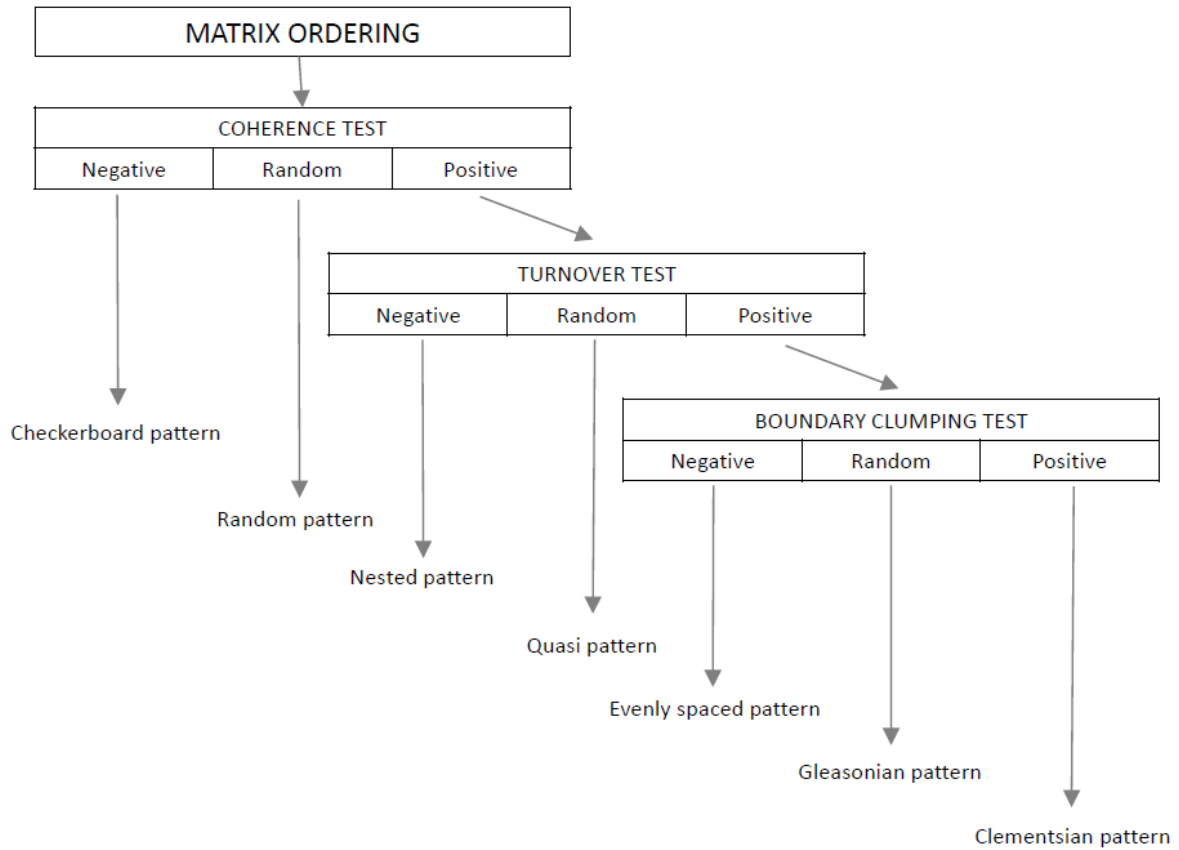
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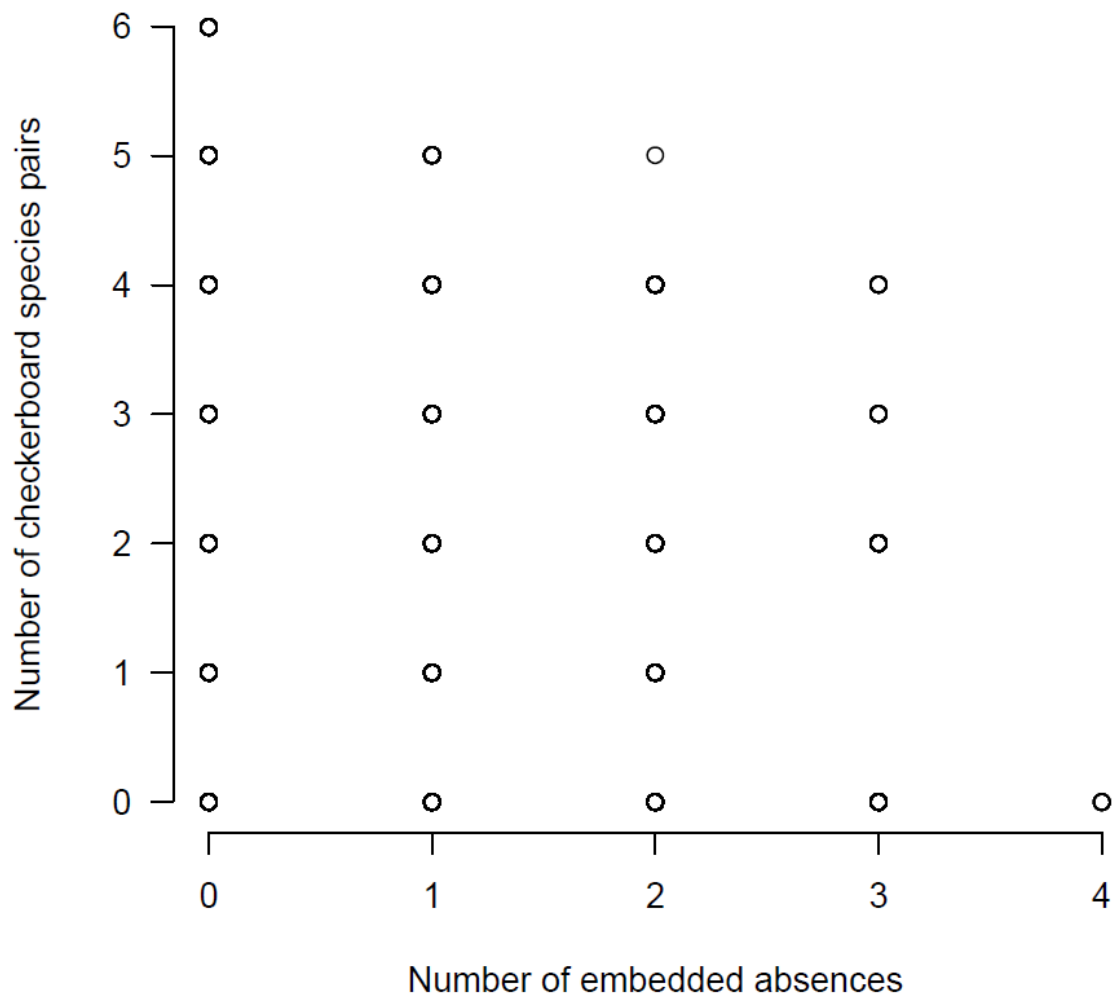
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- 607
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- 609 FIGURES
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611

612 Fig. 1: Diagrammatic representation of the Elements of Metacommunity Structure (EMS)
 613 framework following Leibold and Mikkelsen (2002) and Presley et al. (2010).

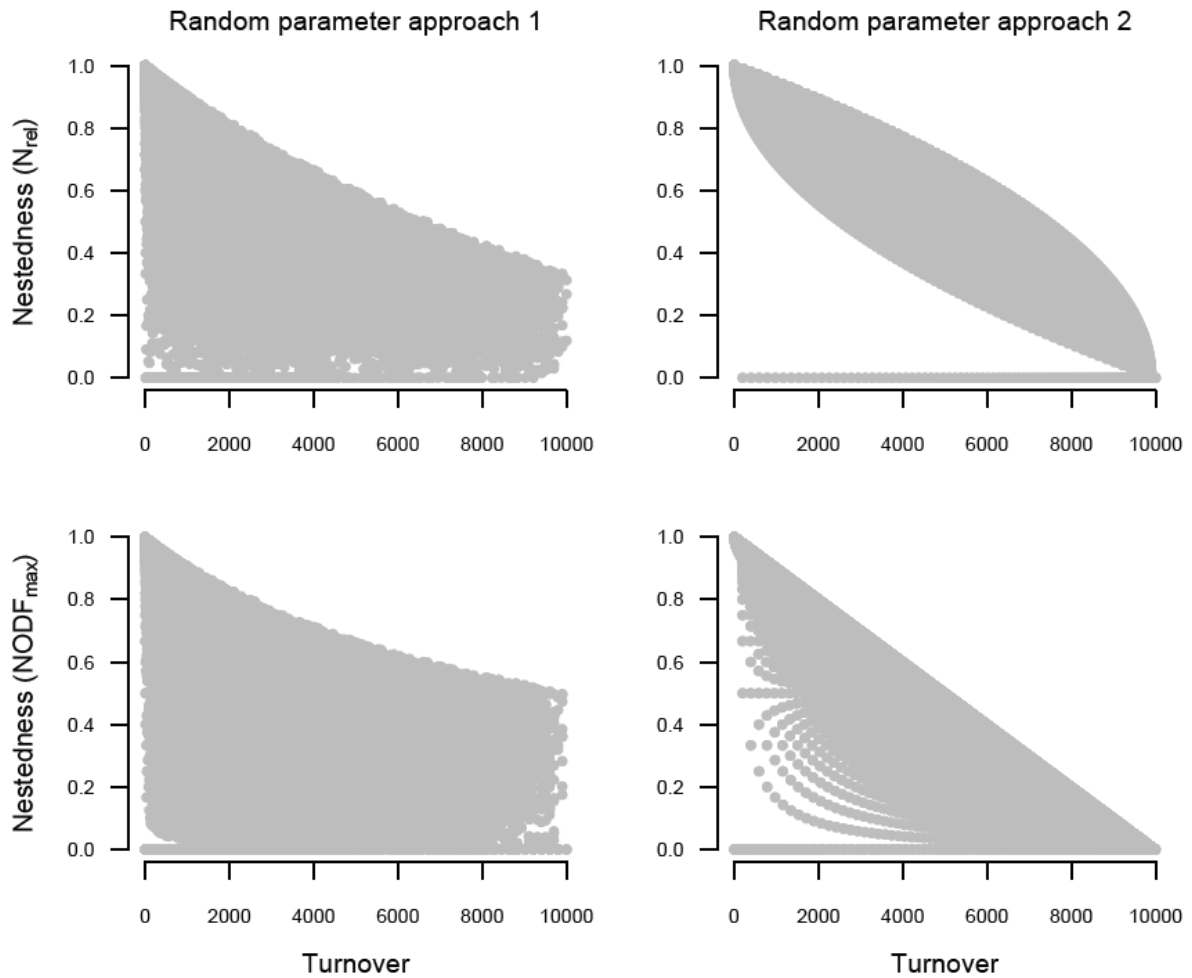
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616 Fig. 2: The relationship of the number of embedded absences and the number of
617 checkerboard species pairs when incidence matrices with 4 sites and 4 species were
618 examined.

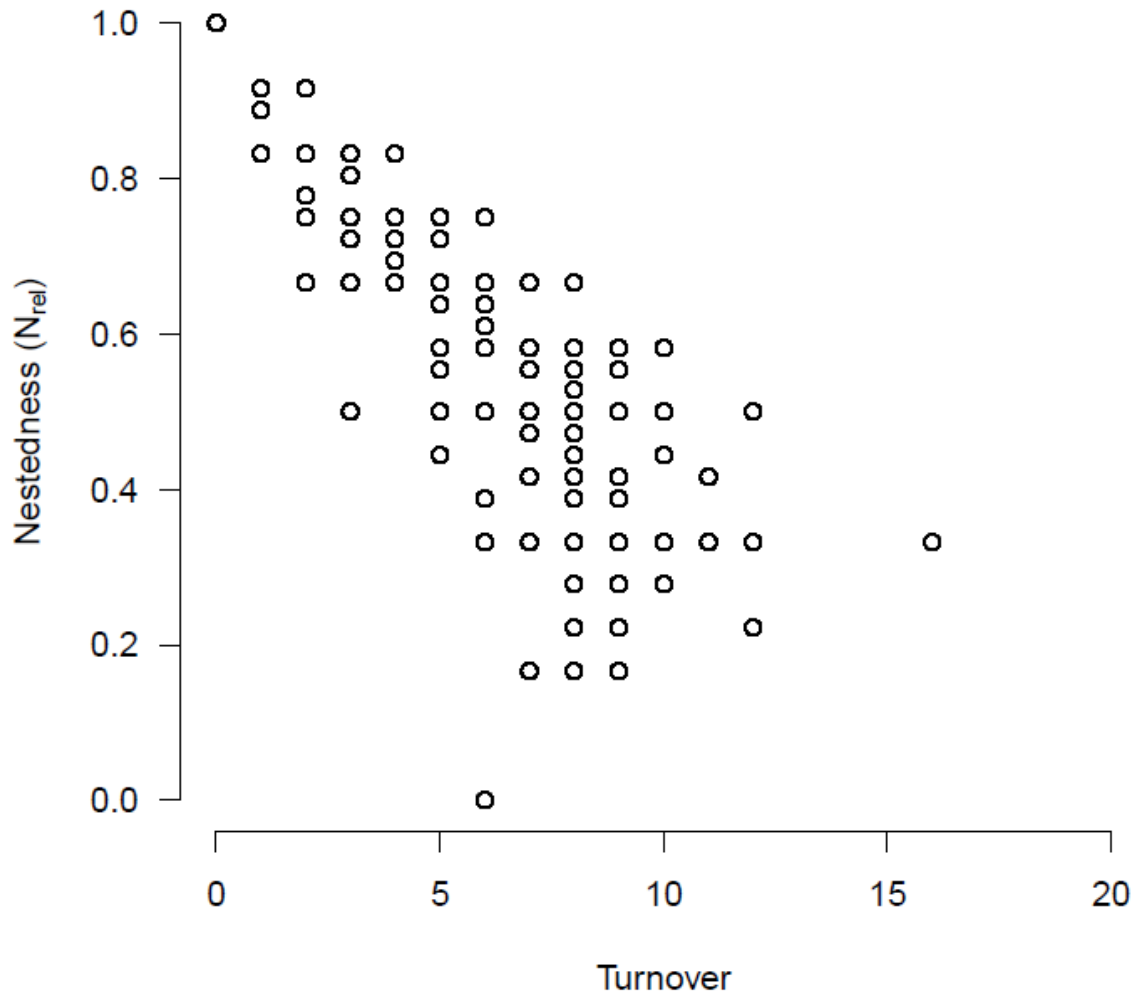
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620

621 Fig. 3: The relationship between turnover (horizontal axes) and nestedness (vertical axes)
 622 when pairs of sites were examined. Upper subfigures show when nestedness was quantified
 623 as N_{rel} , while lower subfigures show when nestedness was quantified as $NODF_{max}$. Left
 624 subfigures show the results of the Random parameter approach 1, while right subfigures
 625 those of the Random parameter approach 2.

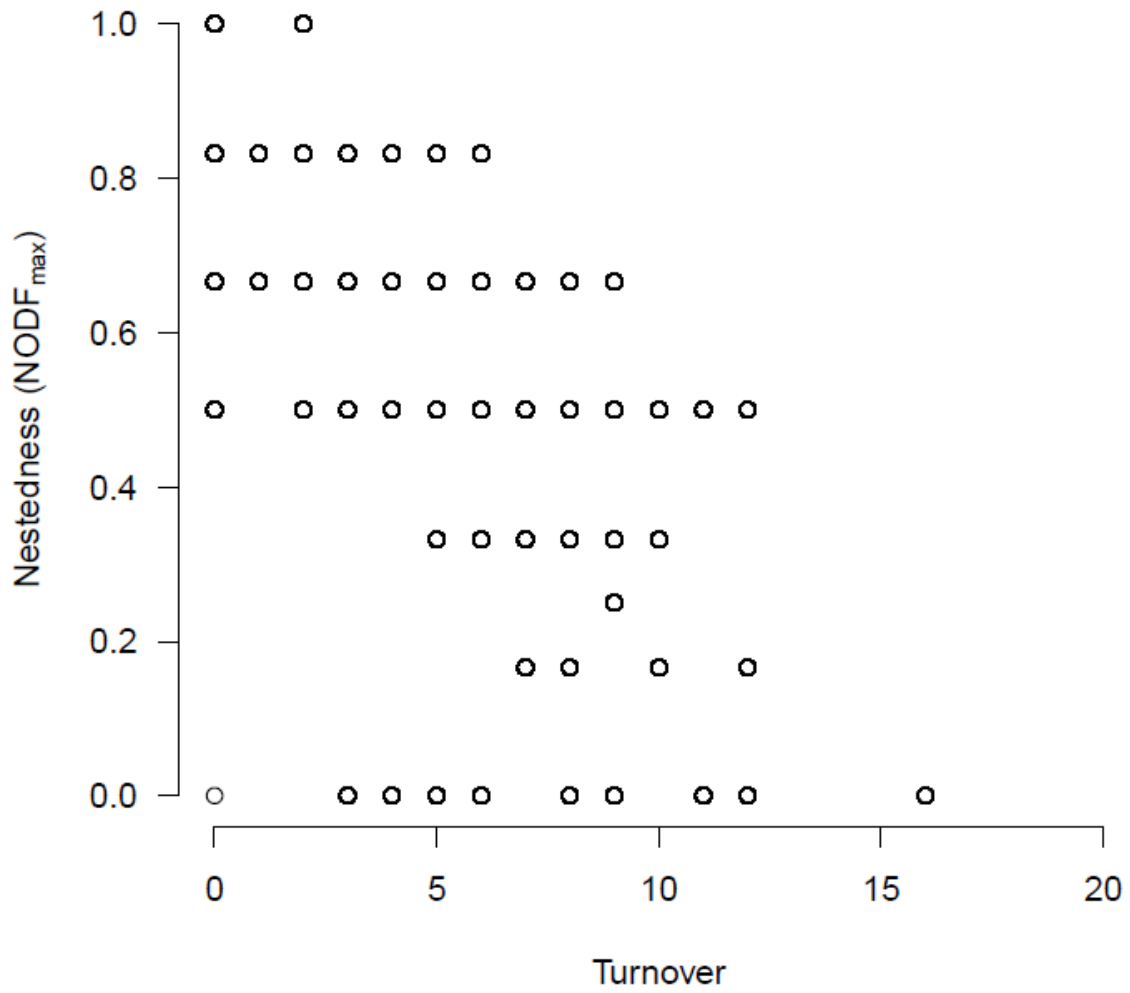
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628 Fig. 4: The relationship between turnover and nestedness (N_{rel}) when incidence matrices
 629 with 4 sites and 4 species were examined.

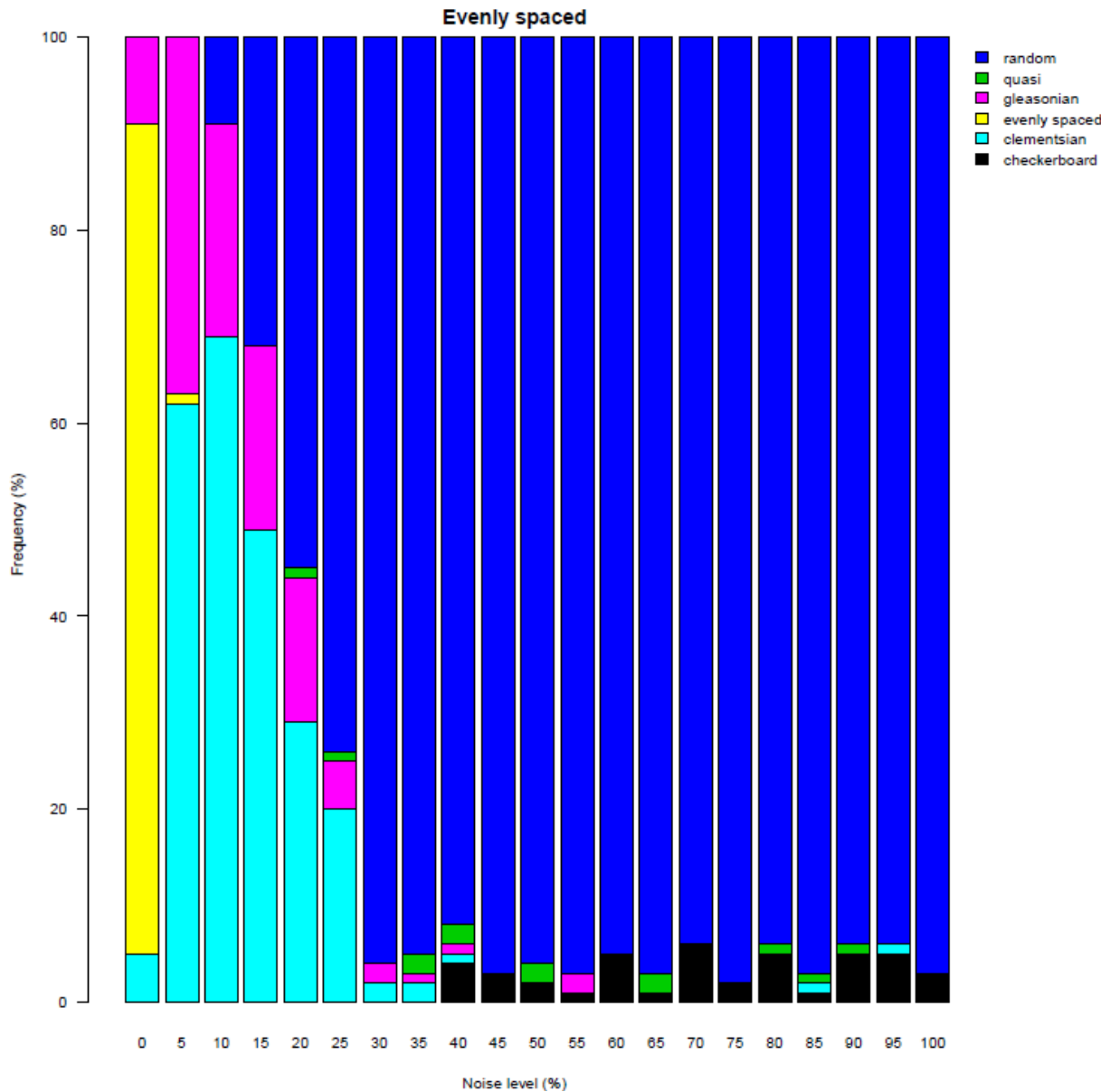
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631

632 Fig. 5: Relationship between turnover and nestedness (NODF_{max}) when incidence matrices
 633 with 4 sites and 4 species were examined.

634



635

636 Fig. 6: Bar plot showing the frequency of idealized metacommunity patterns (vertical axis)
 637 detected by the Elements of Metacommunity Structure (EMS) framework when evenly
 638 spaced pattern was exposed to increasing noise (horizontal axis).

639

640

641

642

643 THE MANUSCRIPT CONTAINS ALSO THE FOLLOWING ELECTRONIC APPENDICES

644

645 Electronic Appendix 1: Idealized metacommunity patterns used in the noise test.

646

647 Electronic Appendix 2: R script used for calculating indices.

648

649 Electronic Appendix 3: The ordering of an incidence matrix with sites with single and unique
650 species. R scripts.

651

652 Electronic Appendix 4: Visualization of 4-by-4 incidence matrices with the largest number of
653 checkerboard species pairs.

654

655 Electronic Appendix 5: The results of the noise tests on nested, Gleasonian and Clementsian patterns.

656